Understanding and evaluating blind deconvolution algorithms

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Abstract

Blind deconvolution is the recovery of a sharp version of a blurred image when the blur kernel is unknown. Recent algorithms have afforded dramatic progress, yet many aspects of the problem remain challenging and hard to understand. The goal of this paper is to analyze and evaluate recent blind deconvolution algorithms both theoretically and experimentally. We explain the previously reported failure of the naive MAP approach by demonstrating that it mostly favors no-blur explanations. On the other hand we show that since the kernel size is often smaller than the image size a MAP estimation of the kernel alone can be well constrained and accurately recover the true blur.

The plethora of recent deconvolution techniques makes an experimental evaluation on ground-truth data important. We have collected blur data with ground truth and compared recent algorithms under equal settings. Additionally, our data demonstrates that the shift-invariant blur assumption made by most algorithms is often violated.

1. Introduction

Blind deconvolution is the problem of recovering a sharp version of an input blurry image when the blur kernel is unknown [10]. Mathematically, we wish to decompose a blurred image \( y \) as

\[
y = k \otimes x
\]

where \( x \) is a visually plausible sharp image, and \( k \) is a non negative blur kernel, whose support is small compared to the image size. This problem is severely ill-posed and there is an infinite set of pairs \((x, k)\) explaining any observed \( y \). For example, One undesirable solution that perfectly satisfies eq. 1 is the no-blur explanation: \( k \) is the delta (identity) kernel and \( x = y \). The ill-posed nature of the problem implies that additional assumptions on \( x \) or \( k \) must be introduced.

Blind deconvolution is the subject of numerous papers in the signal and image processing literature, to name a few consider [1, 8, 22, 15, 17] and the survey in [10]. Despite the exhaustive research, results on real world images are rarely produced. Recent algorithms have proposed to address the ill-posedness of blind deconvolution by characterizing \( x \) using natural image statistics [16, 4, 14, 6, 7, 3, 20]. While this principle has lead to tremendous progress, the results are still far from perfect. Blind deconvolution algorithms exhibit some common building principles, and vary in others. The goal of this paper is to analyze the problem and shed new light on recent algorithms. What are the key challenges and what are the important components that make blind deconvolution possible? Additionally, which aspects of the problem should attract further research efforts?

One of the puzzling aspects of blind deconvolution is the failure of the MAP approach. Recent papers emphasize the usage of a sparse derivative prior to favor sharp images. However, a direct application of this principle has not yielded the expected results and all algorithms have required additional components, such as marginalization across all possible images [16, 4, 14], spatially-varying terms [7, 19], or solvers that vary their optimization energy over time [19]. In this paper we analyze the source of the MAP failure. We show that counter-intuitively, the most favorable solution under a sparse prior is usually a blurry image and not a sharp one. Thus, the global optimum of the MAP approach is the no-blur explanation. We discuss solutions to the problem and analyze the answers provided by existing algorithms. We show that one key property making blind deconvolution possible is the strong asymmetry between the dimensionalities of \( x \) and \( k \). While the number of unknowns in \( x \) increases with image size, the dimensionality of \( k \) remains small. Therefore, while a simultaneous MAP estimation of both \( x \) and \( k \) fails, a MAP estimation of \( k \) alone (marginalizing over \( x \)), is well constrained and recovers an accurate kernel. We suggest that while the sparse prior is helpful, the key component making blind deconvolution possible is not the choice of prior, but the thoughtful choice of estimator. Furthermore, we show that with a proper estimation rule, blind deconvolution can be performed even with a weak Gaussian prior.

Finally, we collect motion-blurred data with ground truth. This data allows us to quantitatively compare recent blind deconvolution algorithms. Our evaluation suggest that the variational Bayes approach of [4] outperforms all existing alternatives. This data also shows that the shift invariance convolution model involved in most existing algorithms is often violated and that realistic camera shake includes in-plane rotations.

2. MAP\(_{x,k}\) estimation and its limitations

In this paper \( y \) denotes an observed blurry image, which is a convolution of an unknown sharp image \( x \) with an unknown blur kernel \( k \), plus noise \( n \) (this paper assumes i.i.d. Gaussian noise):

\[
y = k \otimes x + n.
\]
Using capital letters for the Fourier transform of a signal:

\[ Y_\omega = K_\omega X_\omega + N_\omega. \]  

The goal of blind deconvolution is to infer both \( k \) and \( x \) given a single input \( y \). Additionally, \( k \) is non-negative, and its support is often small compared to the image size.

The simplest approach is a maximum-a-posteriori (MAP) estimation, seeking a pair \( (\hat{x}, \hat{k}) \) maximizing:

\[ p(x, k|y) \propto p(y|x, k)p(x)p(k). \]  

For simplicity of the exposition, we assume a uniform prior on \( k \). The likelihood term \( p(y|x, k) \) is the data fitting term \( \log p(y|x, k) = -\lambda \|k \ast x - y\|^2 \). The prior \( p(x) \) favors natural images, usually based on the observation that their gradient distribution is sparse. A common measure is

\[ \log p(x) = -\sum_k |g_{x, i}(x)|^\alpha + |g_{y, i}(x)|^\alpha + C \]  

where \( g_{x, i}(x) \) and \( g_{y, i}(x) \) denote the horizontal and vertical derivatives at pixel \( i \) (we use the simple \([-1 1]\) filter) and \( C \) is a constant normalization term. Exponent values \( \alpha < 1 \) lead to sparse priors and natural images usually correspond to \( \alpha \) in the range \([0.5, 0.8] \) [21]. Other choices include a Laplacian prior \( \alpha = 1 \), and a Gaussian prior \( \alpha = 2 \). While natural image gradients are very non-Gaussian, we examine this model because it enables an analytical treatment.

The MAP approach seeks \( (\hat{x}, \hat{k}) \) minimizing

\[ (\hat{x}, \hat{k}) = \arg\min_{x,k} \lambda \|k \ast x - y\|^2 + \sum_i |g_{x, i}(x)|^\alpha + |g_{y, i}(x)|^\alpha. \]  

Eq. (6) reveals an immediate limitation:

**Claim 1** Let \( x \) be an arbitrarily large image sampled from the prior \( p(x) \), and \( y = k \ast x \). The pair \((x, k)\) optimizing the MAP score satisfies \( |x| \to 0 \) and \( |k| \to \infty \).

**Proof:** For every pair \((x, k)\) we use a scalar \( s \) to define a new pair \( x' = s \cdot x \), \( k' = 1/s \cdot k \) with equal data fitting \( \|k' \ast x' - y\|^2 = \|k' \ast x' - y\|^2 \). While the data fitting term is constant, the prior term improves as \( s \to 0 \).

This observation is not surprising. The most likely image under the prior in Eq. (5) is a flat image with no gradients. One attempt to fix the problem is to assume the mean intensity of the blurred and sharp images should be equal, and constrain the sum of \( k \): \( \sum_k k_i = 1 \). This eliminates the zero solution, but usually the no-blur solution is still favored.

To understand this, consider the 1D signals in Fig. 1 that were convolved with a (truncated) Gaussian kernel \( k^\star \) of standard deviation 4 pixels. We compare two interpretations: 1) the true kernel: \( y = k^\star \ast x \). 2) the delta kernel (no blur) \( y = k^0 \ast x \). We evaluate the \( -\log p(x, k|y) \) score (Eq. (6)), while varying the \( \alpha \) parameter in the prior.

For step edges (Fig. 1(a)) MAP usually succeeds. The edge is sharper than its blurred version and while the Gaussian prior favors the blurry explanation, appropriate sparse priors (\( \alpha < 1 \)) favor the correct sharp explanation.

In contrast, Fig. 1(b) presents a narrow peak. Blurring reduces the peak height, and as a result, the Laplacian prior \( \alpha = 1 \) favors the blurry \( x \) (\( k \) is delta) because the absolute sum of gradients is lower. Examining Fig. 1(b-right) suggests that the blurred explanation is winning for smaller \( \alpha \) values as well. The sharp explanation is favored only for low alpha values, approaching a binary penalty. However, the sparse models describing natural images are not binary, they are usually in the range \( \alpha \in [0.5, 0.8] \) [21]. The last signal considered in Fig. 1(c) is a row cropped from a natural image, illustrating that natural images contain a lot of medium contrast texture and noise, corresponding to the narrow peak structure. This dominates the statistics more than step edges. As a result, blurring a natural image reduces the overall contrast and, as in Fig. 1(b), even sparse priors favor the blurry \( x \) explanation.
ever, for most natural images the second effect is stronger and blur reduces the average gradient magnitude. (b) Expected negative likelihood reduces probability increases with blur.

To confirm the above observation, we blurred the image in Fig. 2 with a Gaussian kernel of standard deviation 3 pixels. We compared the sum of the gradients in the blurred and sharp images using α = 0.5. For 15 × 15 windows the blurred image is favored over 97% of the windows, and this phenomenon increases with window size. For 45 × 45 windows, the blurred version is favored at all windows. Another observation is that if the sharp explanation does win, it happens next to significant edges.

To understand this, note that blur has two opposite effects on the image likelihood: 1) it makes the signal derivatives less sparse, and that reduces the likelihood. 2) It reduces the derivatives variance and that increases its likelihood. For very specific images, like ideal step edges, the likelihood is more sparse, and that reduces the likelihood. 2) It reduces the derivatives variance and that increases its likelihood. For very specific images, like ideal step edges, the likelihood is more sparse, and that reduces the likelihood.

To see the role of marginalization, consider the scalar blind deconvolution strategy: use a MAP estimator to recover the kernel and, given that, sample the full volume of $p(x, k)$.

Revisiting the literature on the subject, Fergus et al. [4] report that their initial attempts to approach blind deconvolution with MAP$_{x,k}$ failed, resulting in either the original blurred explanation or a binary two-tone image, depending on parameter tunings.

Algorithms like [7, 6] explicitly detect edges in the image (either manually or automatically), and seek a kernel which transfers these edges into binary ones. This is motivated by the example in Fig. 2, suggesting that MAP$_{x,k}$ could do the right thing around step edges. Another algorithm which makes usage of this property is [19]. It optimizes a semi-MAP$_{x,k}$ score, but explicitly detects smooth image regions and reweights their contribution. Thus, the MAP$_{x,k}$ score is dominated by edges. We discuss this algorithm in detail in [13].
Let \( x \) be an arbitrarily large image, sampled from the prior \( p(x) \), and \( y = k \otimes x + n \). Then \( p(k|y) \) is maximized by the true kernel \( k^* \). Moreover, if \( \arg \max_k p(y|k) \) is unique, \( p(k|y) \) approaches a delta function\(^2\).

\(^2\)Note that Claim 2 does not guarantee that the MAP\(_k\) estimate is unique. For example, if the kernel support is not constrained enough, multiple spatial shifts of the kernel provide equally good solutions. The problem can be easily avoided by a weak prior on \( k \) (e.g. favoring centered kernels).

**Proof:** We divide the image into small disjoint windows \( \{y^1, ..., y^n\} \) and treat them as i.i.d. samples \( y^j \sim p(y|k^*) \). We then select \( k^{ML} = \arg \max_k \prod_j p(y^j|k) \). Applying the standard consistency theorem for maximum likelihood estimators [9] we know that given enough samples, the ML approachs the true parameters. That is, when \( n \to \infty \)

\[
p(k^{ML}(\{y^1, ..., y^n\}) = k^*) \to 1.
\] (10)

Due to the local form of the prior \( p(x) \) (Eq. (5)), taking sufficiently far away disjoint windows will ensure that \( p(y|k) \approx \prod_j p(y^j|k) \). Thus, \( p(y|k) \) is maximized by \( k^{ML} \).

Also, if we select a \( m \) times larger image \( y' \), \( p(y'|k) = p(y|k)^m \). Thus, if \( p(y|k) < \max_k p(y|k) \) then \( p(y|k) \to 0 \).

Finally, if \( p(k^*) > 0 \), then \( k^{MAP}, k^{ML} \) are equal on large images since \( \arg \max_k p(y|k) = \arg \max_k p(y|k)p(k) \), and thus, \( k^{MAP} \to k^* \). Similarly, if \( \max_k p(y|k) \) is unique, \( p(k|y) \) approaches a delta function.

Fig. 4(c) plots \( p(y|k) \) for a scalar blind deconvolution task with \( N \) observations \( y_j = kx_j + n_j \), illustrating that as \( N \) increases, the uncertainty around the solution decreases (compare with Fig. 4(b)).

In [13] we also justify the MAP\(_k\) approach from the loss function perspective.

### 3.1. Examples of MAP\(_k\) estimation

Claim 2 reduces to a robust blind deconvolution strategy: use MAP\(_k\) estimator to recover \( k^{MAP} = \arg \max_k p(k|y) \), and then use \( k^{MAP} \) to solve for \( x \) using some non blind deconvolution algorithm. To illustrate the MAP\(_k\) approach, we start with the simple case of a Gaussian prior on \( p(x) \), as it permits a derivation in closed form.

#### 3.1.1 The Gaussian prior

The prior on \( X \) in Eq. (5) is a convolution and thus diagonal in the frequency domain. If \( G_x, G_y \) denote the Fourier transform of the derivatives \( g_x, g_y \), then:

\[
X \sim N(0, diag(\sigma_x^2)) \quad \sigma_x^2 = \beta(\|G_x, \omega\|^2 + \|G_y, \omega\|^2)^{-1}.
\] (11)

Note that since a derivative filter is zero at low frequencies and high at higher frequencies, this is similar to the classical 1/f\(^2\) power spectrum law for images. Denoting noise variance by \( \eta \), we can express \( p(X, Y; K) = p(Y|X; K)p(X) \) as:

\[
p(X, Y; K) \propto e^{-\frac{1}{2\sigma_x^2}\|K \cdot X - Y \cdot \omega\|^2 - \frac{\eta}{\sigma_x^2}\|X\cdot \omega\|^2}.
\] (12)

(see [13] for details). Conditioned on \( k \), the mean and mode of a Gaussian are equal:

\[
X^{MAP} = \left( |K|^{-2} + \frac{\eta^2}{\sigma_x^2} \right)^{-1} K^* Y \omega.
\] (13)

Eq. (13) is the classic Wiener filter [5]. One can also integrate \( X \) and express \( p(Y|K) \) analytically. This is also a diagonal zero mean Gaussian with

\[
Y \sim N(0, diag(\phi_\omega^2)) \quad \phi_\omega^2 = \sigma_x^2 |K \cdot \omega|^2 + \eta^2.
\] (14)
Eq. (14) is maximized when \( \phi^2 = |Y_\omega|^2 \), and for blind deconvolution, this implies:

\[
|\hat{K}_\omega|^2 = \max \left( 0, \frac{|Y_\omega|^2 - \eta^2}{\sigma^2} \right).
\]  

(15)

The image estimated using \( \hat{K} \) satisfies \( |X_\omega|^2 = \sigma^2 \). Therefore MAP\( \kappa \) does not result in a trivial \( X = 0 \) solution as MAP\( x, k \) would, but in a solution whose variance matches the prior variance \( \sigma^2 \), that is, a solution which looks like a typical sample from the prior \( p(X) \).

Another way to interpret the MAP\( x, k \), is to note that

\[
\log p(Y|K) = \log p(X^{MAP}, Y; K) - \frac{1}{2} \sum_\omega \log \left( \frac{|K_\omega|^2}{\eta^2} + \frac{1}{\sigma^2} \right) + C
\]

(16)

Referring to Eq. (12), the second term is just the log determinant of the covariance of \( p(X|Y; K) \). This second term is optimized when \( K_\omega = 0 \), i.e. by kernels with more blur. That is, \( \log p(Y|K) \) is equal to the MAP\( x, k \) score of the mode plus a term favoring kernels with blur.

The discussion above suggests that the Gaussian MAP\( x, k \) provides a reasonable solution to blind deconvolution. In the experiment section we evaluate this algorithm and show that, while weaker than the sparse prior, it can provide acceptable solutions. This stands in contrast to the complete failure of a MAP\( x, k \) approach, even with the seemingly better sparse prior. This demonstrates that a careful choice of estimator is actually more critical than the choice of prior.

Note that Eq. (15) is accurate if every frequency is estimated independently. In practice, the solution can be further constrained, because the limited spatial support of \( \hat{K} \) implies that the frequency coefficients \( \{K_\omega\} \) are linearly dependent. Another important issue is that Eq. (15) provides information on the kernel power spectrum alone but leaves uncertainty about the phase. Many variants of Gaussian blind deconvolution algorithms are available in the image processing literature (e.g. [8, 15]) but in most cases only symmetric kernels are considered since their phase is known to be zero. However, realistic camera shake kernels are usually not symmetric. In [13] we describe a Gaussian blind deconvolution algorithm which attempts to recover non symmetric kernels as well.

### 3.1.2 Approximation strategies with a sparse prior

The challenge with the MAP\( x, k \) approach is that for a general sparse prior, \( p(k|y) \) (Eq. (7)) cannot be computed in closed form. Several previous blind deconvolution algorithms can be viewed as approximation strategies for MAP\( x, k \), although the authors might not have motivated them in this way.

A simple approximation is proposed by Levin [14], for the 1D blur case. It assumes that the observed derivatives of \( y \) are independent (this is usually weaker than assuming independent derivatives of \( x \)): \( \log p(y|k) = \sum_i \log p(g_{x,i}(y)|k) \). Since \( p(g_{x,i}(y)|k) \) is a 1D distribution, it can be expressed as a 1D table, or a histogram \( h^k \).

The independence assumption implies that instead of summing over image pixels, one can express \( p(y|k) \) by summing over histogram bins:

\[
\log p(y|k) = \sum_i \log p(g_{x,i}(y)|k) = \sum_j h_j \log(h_j^k)
\]

(17)

where \( h \) denotes the gradients histogram in the observed image and \( j \) is a bin index. In a second step, note that maximizing Eq. (17) is equivalent to minimizing the histogram distance between the observed and expected histograms \( h, h^k \). This is because the Kullback Leibler divergence is equal to the negative log likelihood, plus a constant that does not depend on \( k \) (the negative entropy):

\[
D_{KL}(h, h^k) = \sum_j h_j \log(h_j^k) - \sum_j h_j \log(h_j).
\]

(18)

Since the KL divergence is non-negative, the likelihood is maximized when the histograms \( h, h^k \) are equal. This very simple approach is already able to avoid the delta solution but as we demonstrate in Sec. 4.1 it is not accurately identifying the exact filter width.

A stronger approximation is the variational Bayes meanfield approach taken by Fergus et al. [4]. The idea is to build an approximating distribution with a simpler parametric form:

\[
p(x, k|y) \approx q(x, k) = \prod_i q(g_{x,i}(x)) q(g_{y,i}(x)) \prod_j q(k_j).
\]

(19)

Since \( q \) is expressed in the gradient domain this does not recover \( x \) directly. Thus, they also pick the MAP\( x \) kernel from \( q \) and then solve for \( x \) using non blind deconvolution.

A third way to approximate the MAP\( x, k \) is the Laplace approximation [2], which is a generalization of Eq. (16):

\[
\log p(y|k) \approx \log p(x^{MAP}, y; k) - \frac{1}{2} \log |A| + C
\]

(20)

\[
A = \frac{\partial^2}{\partial x_i \partial x_j} \log p(x, y; k)_{x = x^{MAP}}.
\]

(21)

The Laplace approximation states that \( p(y|k) \) can be expressed by the probability of the mode \( x^{MAP} \) plus the log determinant of the variance around the mode. As discussed above, higher variance is usually achieved when \( k \) contains more zero frequencies, i.e. more blur. Therefore, the Laplace approximation suggests that \( p(y|k) \) is the MAP\( x, k \) score plus a term pulling toward kernels with more blur. Unfortunately, in the non Gaussian case the covariance matrix isn’t diagonal and exact inversion is less trivial. Some earlier blind deconvolution approaches [22, 17] can be viewed as simplified forms of a blur favoring term. For example, they bias towered blurry kernels by adding a term penalizing the high frequencies of \( k \) or with an explicit prior on the kernel. Another approach was exploit by Bronstein et al. [3]. They note that in the absence of noise and with invertible kernels \( p(k|y) \) can be exactly evaluated for sparse priors as well. This reduces to optimizing the sparsity of the image plus the log determinant of the kernel spectrum.
4. Evaluating blind deconvolution algorithms

In this section we qualitatively compare blind deconvolution strategies on the same data. We start with a synthetic 1D example and in the second part turn to real 2D motion.

4.1. 1D evaluation

As a first test, we use a set of 1000 signals of size 10 × 1 cropped from a natural image. These small 1D signals allow us to evaluate the marginalization integral in Eq. (7) exactly even for a sparse prior. The signals were convolved with a 5-tap box filter (cyclic convolution was used) and an i.i.d. Gaussian noise with standard deviation 0.01 was added. We explicitly search over the explanations of all box filters of size ℓ = 1, ..., 7 taps (all filters normalized to 1). The explicit search allows comparison of the score of different blind deconvolution strategies without folding in optimization errors. (In practice optimization errors do have a large effect on the successes of blind deconvolution algorithms.)

The exact −log p(y|k) score is minimized by the true box width ℓ = 5.

We tested the zero sheet separation (e.g. [11]), an earlier image processing approach with no probabilistic formulation. This algorithm measures the Fourier magnitude of y at the zero frequencies of each box filter k. If the image was indeed convolved with that filter, low Fourier content is expected. However, this approach considers the zero frequencies alone ignoring all other information, and is known to be noise sensitive. It is also limited to kernel families from a simple parametric form and with a clear zeros structure.

Supporting the example in Sec. 2, a pure MAP_x,k approach (p(y|k) ≈ p(x|MAP, y|k)) favors no-blur (ℓ = 1). Reweighting the derivative penalty around edges can improve the situation, but the delta solution still provides a noticeable local optimum.

The correct minimum is favored with a variational Bayes approximation [4] and with the semi Laplace approximation of [3]. The independence approximation [14] is able to overcome the delta solution, but does not localize the solution very accurately (minimum at ℓ = 4 instead of ℓ = 5.) Finally, the correct solution is identified even with the poor image prior provided by a Gaussian model, demonstrating that the choice of estimator (MAP_x,k v.s. MAP_k), is more critical than the actual prior (Gaussian v.s. sparse).

Since claim 2 guarantees success only for large images, we attempt to evaluate how large an image should be in practice. Fig. 6 plots the uncertainty in p(k|y) for multiple random samples of N = 10 × 1 columns. The probability is tightly peaked at the right answer for as little as N = 20 columns. The search space in Fig. 6 is limited to the single parameter family of box filters. In real motion deblurring one searches over a larger family of kernels and a larger uncertainty is expected.

4.2. 2D evaluation

To compare blind deconvolution algorithms we have collected blurred data with ground truth. We capture a sharp version a planar scene (Fig. 7(a)) by mounting the camera on a tripod, as well as a few blurred shots. Using the sharp reference we solve for a non-negative kernel k minimizing ∥k ⊗ x − y∥^2. The scene in Fig. 7(a) includes high frequency noise patterns which helps stabilizing the constraints on k.

The central area of the frame includes four real images used as input to the various blind deconvolution algorithms.

We first observed that assuming a uniform blur over the image is not realistic even for planar scenes. For example Fig. 7(b) shows traces of points at 4 corners of an image captured by a hand-held camera, with a clear variation between the corners. This suggests that an in-plane rotation (rotation around the z-axis) is a significant component of human hand shake. Yet, since a uniform assumption is made by most algorithms, we need to evaluate them on data that is realistic.
We used an 85mm lens and a 0.3 seconds exposure. The kernels’ support varied from 10 to 25 pixels.

We can measure the SSD error between a deconvolved output and the ground truth. However, wider kernels result in larger deconvolution error even with the true kernel. To normalize this effect, we measure the ratio between deconvolution error with the estimated kernel and deconvolution with the truth kernel. In Fig. 9 we plot the cumulative histogram of error ratios (e.g. bin \( r = 3 \) counts the percentage of test examples achieving error ratio below 3). Empirically, we noticed that error ratios above 2 are already visually implausible. One test image is presented in Fig. 10, all others included in [13].

We have evaluated the algorithms of Fergus et al. [4] and Shan et al. [19] (each using the authors’ implementation), as well as MAP\(_k\) estimation using a Gaussian prior [13], and a simplified MAP\(_{x,k}\) approach constraining \( \sum k_i = 1 \) (we used coordinate descent, iterating between holding \( x \) constant and solving for \( k \), and then holding \( k \) constant and solving for \( x \)). The algorithms of [14, 7, 3] were not tested because the first was designed for 1D motion only, and the others focus on smaller blur kernels.

We made our best attempt to adjust the parameters of Shan et al. [19], but run all test images with equal parameters. Fergus et al. [4] used Richardson-Lucy non blind deconvolution in their code. Since this algorithm is a source for ringing artifacts, we improved the results using the kernel estimated by the authors’ code with the (non blind) sparse deconvolution of [12]. Similarly, we used sparse deconvolution with the kernel estimated by Shan et al. [19].

The bars in Fig. 9 and the visual results in [13] suggest that Fergus et al.’s algorithm [4] significantly outperforms all other alternatives. Many of the artifacts in the results of [4] can be attributed to the Richardson-Lucy artifacts, or to non uniform blur in their test images. Our comparison also suggests that applying sparse deconvolution using the kernels outputted by Shan et al. [19] improves their results. As expected, the naive MAP\(_{x,k}\) approach outputs small kernels approaching the delta solution.

5. Discussion

This paper analyzes the major building blocks of recent blind deconvolution algorithms. We illustrate the limitation of the simple MAP\(_{x,k}\) approach, favoring the no-blur (delta kernel) explanation. One class of solutions involves explicit edge detection. A more principled strategy exploits the dimensionality asymmetry, and estimates MAP\(_k\) while marginalizing over \( x \). While the computational aspects involved with this marginalization are more challenging, existing approximations are powerful.

We have collected motion blur data with ground truth and quantitatively compared existing algorithms. Our comparison suggests that the variational Bayes approximation [4] significantly outperforms all existing alternatives. The conclusions from our analysis are useful for directing future blind deconvolution research. In particular, we
note that modern natural image priors [18, 23] do not overcome the $\text{MAP}_{x,k}$ limitation (and in our tests did not change the observation in Sec. 2). While it is possible that blind deconvolution can benefit from future research on natural image statistics, this paper suggests that better estimators for existing priors may have more impact on future blind deconvolution algorithms. Additionally, we observed that the popular spatially uniform blur assumption is usually unrealistic. Thus, it seems that blur models which can relax this assumption [20] have a high potential to improve blind deconvolution results.

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